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# Bayesian solutions to biodosimetry count data problems, and software solutions

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# Review

Ainsbury EA, Vinnikov VA, Puig P, Higuera M, Maznyk NA, Lloyd DC, Rothkamm K (2014). *Review of Bayesian Statistical Analysis Methods for Cytogenetic Radiation Biodosimetry, with a Practical Example. Radiat. Prot. Dosim.* **162**(3), 185-196.

# The Bayesian approach

A Bayesian approach is highly applicable to ionising radiation dosimetry data. It allows the cytogenetic experts to consider prior knowledge surrounding an overexposure scenario.

This approach implies an accurate measure of the uncertainty of dose estimates.

The calibrative density is the solution of the Bayesian inverse regression problem,

$$P(x|y) \propto P(x) \int L(y|\Theta)P(\Theta|x)d\Theta.$$

# Groer & Pereira (1987)

Poisson responses and just linear dose-response (without intercept).

$$Y_i \sim \text{Pois}(\alpha d_i), \alpha \sim \text{Gamma}(a, b) \implies \alpha|Y \sim \text{Gamma}(a + S, b + N),$$

**S** is the total number of aberrations and

$$N = \sum n_i d_i.$$

The absorbed dose

$$D \sim \text{Gamma}(A, B),$$

and the calibrative dose density results

$$p(D|Y) \propto \frac{D^{A+s-1}}{e^{BD}(nD + a + N)^{S+s+b}}.$$

# Brame & Groer (2002)

NB responses with just linear dose-response (without intercept)

$$Y_i \sim \text{NB}(\alpha d_i, \Psi), \Psi \sim \text{Gamma}, P(\alpha, \Psi) = P(\alpha)P(\Psi)$$

and the priors for  $\alpha$  are Uniform or Normal.

The NB model is compared to the Poisson one using the Bayes Factor

$$BF = \frac{\int L_{\text{NB}}(Y|\alpha, \Psi)P(\alpha)P(\Psi)d\alpha d\Psi}{\int L_{\text{Pois}}(Y|\alpha)P(\alpha)d\alpha}.$$

The calibrative dose density results

$$P(D|Y) \propto P(D) \int \frac{\Gamma(\Psi^{-1} + s)}{\Gamma(s + 1)\Gamma(\Psi^{-1})} \left( \frac{\Psi n \alpha D}{1 + \Psi n \alpha D} \right)^s \left( \frac{1}{1 + \Psi n \alpha D} \right)^{1/\Psi} P(\alpha, \Psi|Y) d\alpha d\Psi.$$

The integrals in this methodology are done using numerical integration.

# Other Bayesian works in biodosimetry

- **Madruga *et al.* 1994, 1996:** log-normal model.
- **Kottas *et al.* 2002:** non-parametric model.
- **Serna *et al.* 2008:** Jeffrey's prior for analyzing the background distribution.

# Whole body exposure

Higuera M, Puig P, Ainsbury EA, Rothkamm K (2015a). A new Inverse Regression Model Applied to Radiation Biodosimetry. *Proc. R. Soc. A*, DOI: 10.1098/rspa.2014.0588.

# Poisson models

The posterior of the population mean is approximated to a normal, by asymptotic normality of the posterior distribution for large samples and the delta-method, i.e.:

$$\mu|D \sim N \left( f(D, \hat{\beta}), \nabla \cdot \hat{\Sigma} \cdot \nabla^T \right).$$

The calibrative density results

$$P(D|Y) \propto P(D)P(X_D = s),$$

where  $X_D$  is Hermite distributed. If  $\mu/D$  is approximated by a gamma,  $X_D$  is NB distributed.

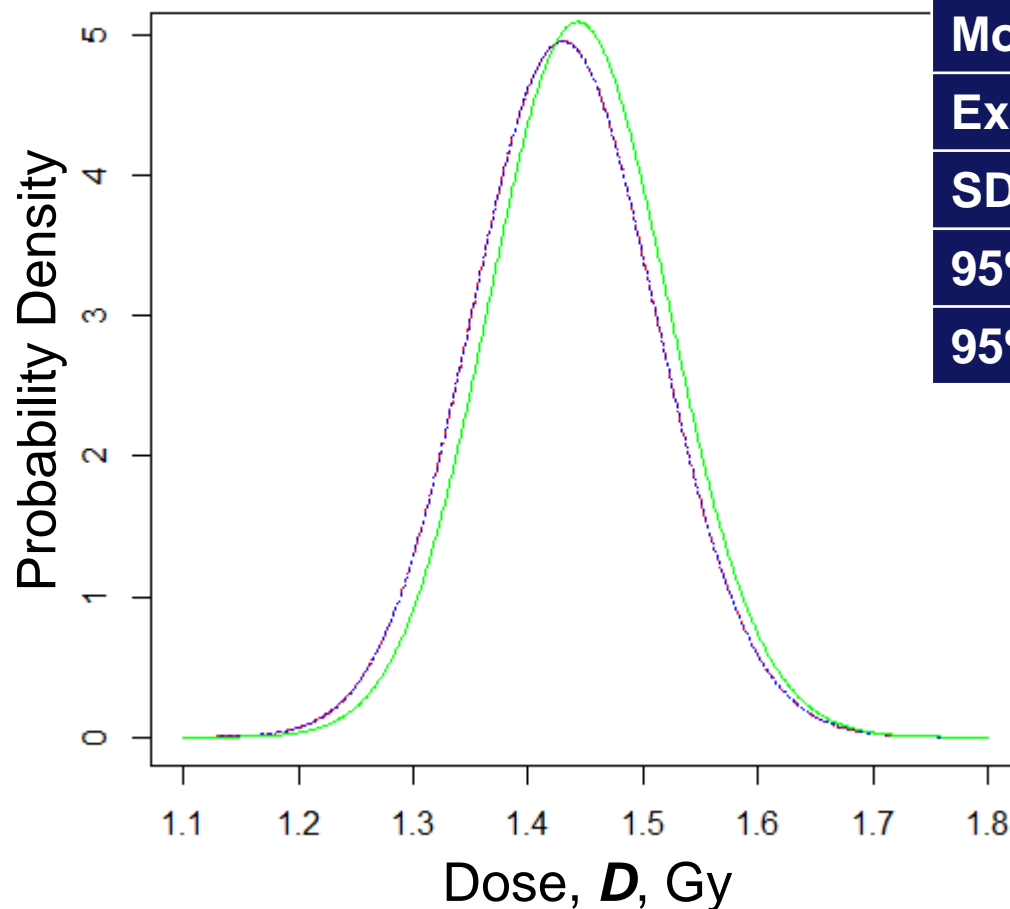


# Example: Romm *et al.* (2013)

- Blood samples from 8 healthy donors were irradiated in vitro with  $^{60}\text{Co}$  gamma-rays at a high-dose rate simulating acute whole body exposure.
- Dicentrics assay: Poisson responses and quadratic dose-response (without intercept).
- The 1.5 Gy sample is removed from the calibration dataset to be used as test data.

Dose (Gy)	Number of dicentrics					$\bar{y}$	$d$	$u$
	0	1	2	3	4			
0.25	2185	8				0.004	0.997	-0.113
0.75	2550	44	1			0.018	1.026	0.952
1.00	2231	54	2			0.025	1.044	1.503
1.50	1712	96	3			0.056	1.003	0.092
2.50	1196	123	7	1		0.105	1.038	0.985
3.00	1070	320	41	6	1	0.295	1.012	0.334

# Example: Romm *et al.* (2013)



Model	(a)	(b)	(c)
Mode	1.430	1.430	1.443
Expected	1.432	1.432	1.445
SD	0.081	0.081	0.078
95% CI LB	1.277	1.277	1.294
95% CI UB	1.594	1.593	1.602

(a): normal mean prior,  
 $U(0, \infty)$  dose prior.

(b): gamma mean prior,  
 $U(0, \infty)$  dose prior.

(c): gamma mean prior,  
 $Ga(21.8, 12.4)$  dose prior.

# Compound Poisson models

The joint posterior of the population mean and the dispersion index is defined:

$$(\mu, \delta) | D \sim N_2 \left( (f(D, \hat{\beta}), \hat{\delta}), \nabla \cdot \hat{\Sigma} \cdot \nabla^T \right).$$

The calibrative density can be defined directly and calculated by numerical integration (not computationally intensive, always bivariate:  $\mathbf{D}$  and  $\boldsymbol{\delta}$ ).

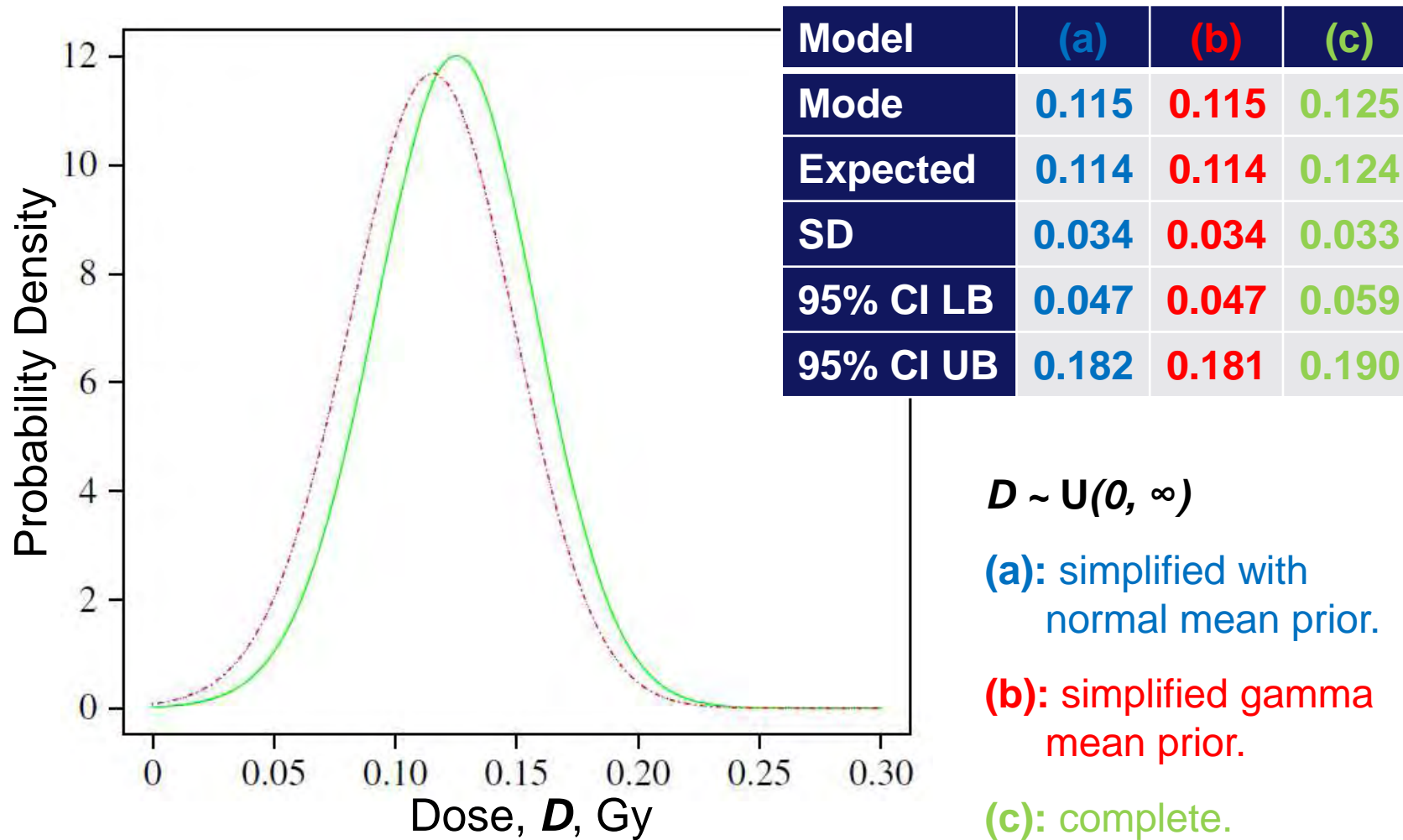
Fixing  $\boldsymbol{\delta}$  by its MLE, the model is reduced and the mean prior is applied like in the Poisson models. The resulting calibrative density is in terms of Compound -Hermite (-NB) mass function.

# Example: Puig & Valero (2006)

- 11 samples of peripheral blood exposed to different doses of  $\gamma$ -rays (0.93 cGymin<sup>-1</sup> dose rate). For each sample, approximately 5000 binucleated cells were inspected.
- MN assay: NB responses and quadratic dose-response.
- The 0.1 Gy sample is used as test data.

dose (Gy)	no. micronuclei									$\bar{y}$	$d$	$u$
	0	1	2	3	4	5	6	7				
0.00	4887	106	5	2						0.024	1.156	7.839
0.10	4773	206	19	2						0.050	1.150	7.526
0.25	4261	324	41	12	2					0.090	1.306	15.306
0.50	4536	364	76	17	7					0.119	1.449	22.484
0.75	4383	512	85	18	2					0.149	1.257	12.876
1.00	4225	636	115	19	5					0.189	1.240	12.009
1.50	4018	805	139	26	9	1	2			0.243	1.270	13.495
2.00	3499	1194	238	45	13	10	1			0.383	1.209	10.471
2.50	3171	1313	393	94	24	3	2			0.501	1.201	10.077
3.00	2582	1575	598	190	44	9	2	6		0.722	1.206	10.307
4.00	1974	1674	869	342	102	26	13	2		1.013	1.172	8.628

# Example: Puig & Valero (2006)



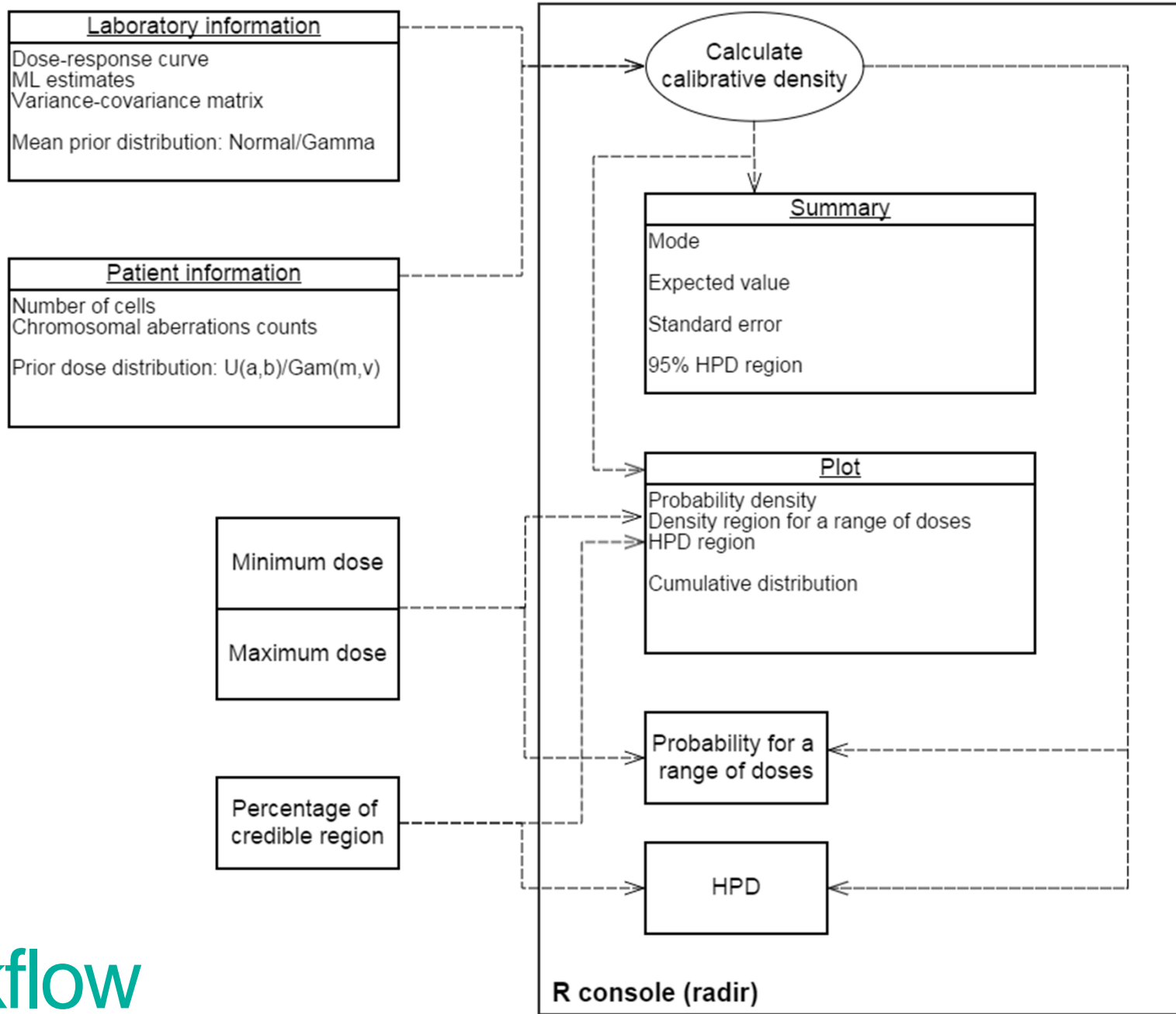
# Conclusions of Higuera *et al.* (2015a)

- New whole body cytogenetic dose estimation Bayesian-like models for Poisson and two parameter compound Poisson responses are presented.
- The approach is valid for any given dose-response function one time differentiable in the parameter set domain.
- To use this methodology, only the estimates of the parameters and covariance matrix of the dose-response curve are required.

# radir

Moriña D, Higuera M, Puig P, Ainsbury EA, Rothkamm K (2015).  
radir package: An R implementation for cytogenetic biodosimetry  
dose estimation. *J. Radiol. Prot.*, DOI: 10.1088/0952-4746/35/3/557.

<https://cran.r-project.org/web/packages/radir/>



# Workflow

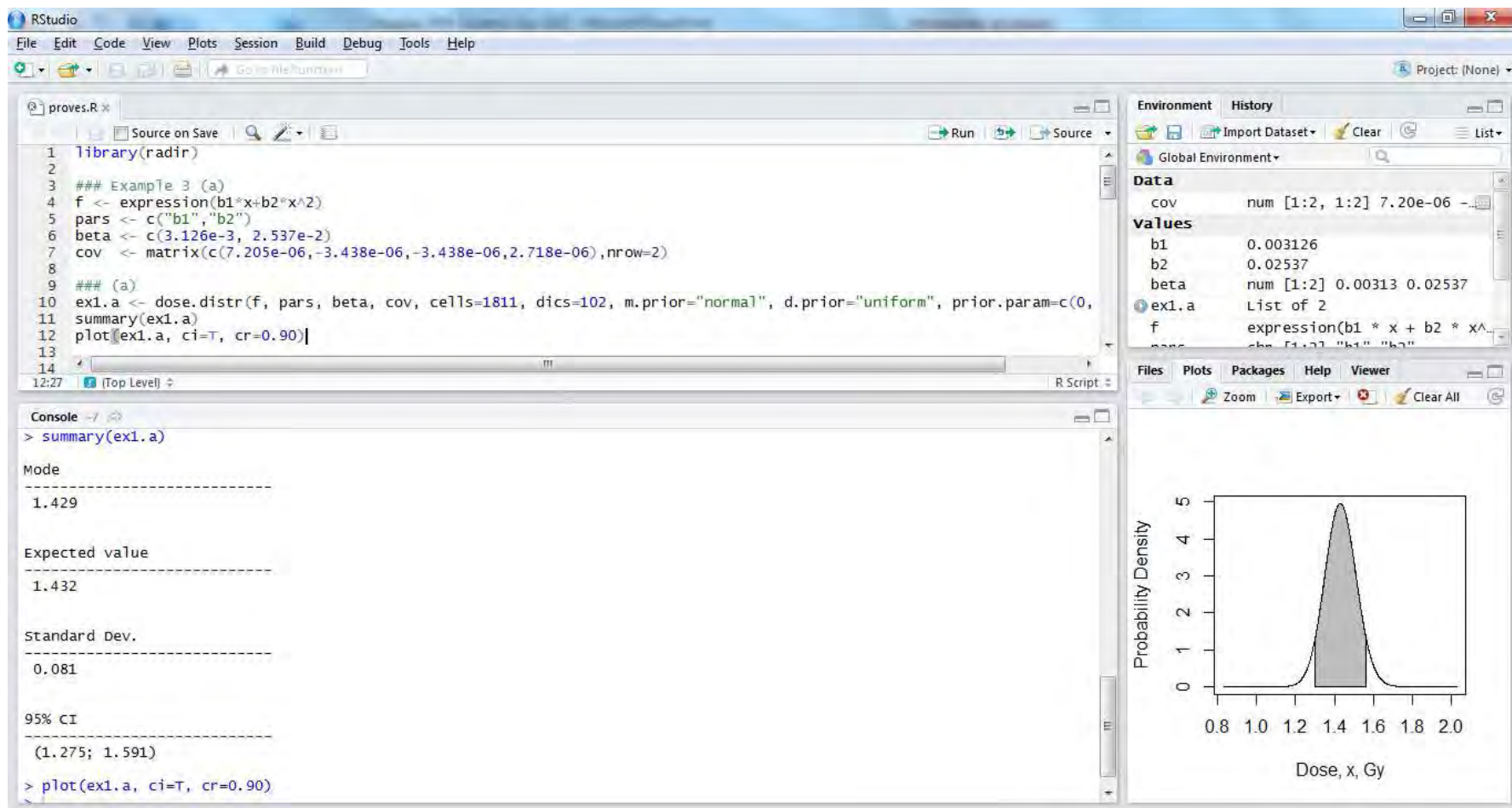


# Video tutorial

[http://polimedia.uab.cat/#v\\_592](http://polimedia.uab.cat/#v_592)

The screenshot displays an RStudio interface with a terminal window on the left and a script editor on the right. The terminal shows several warning messages: "Warning messages: 1: In plot.window(...): 'cr.col' is not a graphical parameter", "2: In plot.xy(xy, type, ...): 'cr.col' is not a graphical parameter", "3: In axis(side = side, at = at, labels = lab): 'cr.col' is not a graphical parameter", "4: In axis(side = side, at = at, labels = lab): 'cr.col' is not a graphical parameter", "5: In box(...): 'cr.col' is not a graphical parameter", "6: In title(...): 'cr.col' is not a graphical parameter". The script editor contains R code for defining parameters, creating dose distributions (ex2.u1, ex2.u2, ex2.u3), and plotting them. A plot is visible at the bottom of the script editor, showing a histogram of 'Dose, x, Gy' with a y-axis labeled '0.1' and an x-axis from 0.0 to 2.0. A man in a plaid shirt is standing behind the RStudio window, and a Polimedia UAB logo is in the top right corner. A video player control bar is at the bottom of the screenshot.

# Screenshot



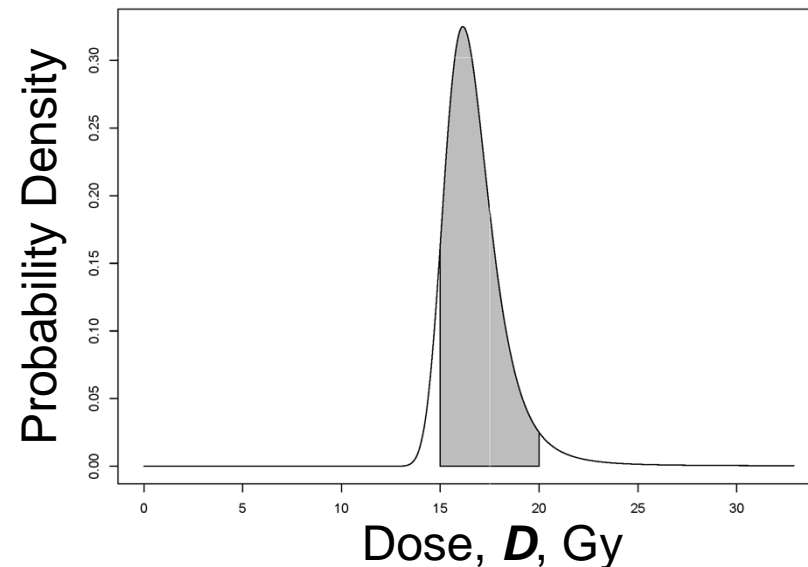
# Example: Pujol *et al.* (2014)

The mean of dicentric per cell as a function of the absorbed dose is

$$f(x; \beta) = \beta_0 e^{-\beta_1 e^{-\beta_2 x}} \left( 1 + \frac{\beta_3 x (2\beta_0 e^{-\beta_1 e^{-\beta_2 x}} + 1)}{1 + \beta_3 x (\beta_0^2 (e^{-\beta_1 e^{-\beta_2 x}})^2 + \beta_0 e^{-\beta_1 e^{-\beta_2 x}})} \right).$$

We assume Poisson responses. The 17 Gy test sample consists in 914 dicentric scored in 100 cells. Assuming the gamma mean prior and a  $U(0, \infty)$  prior dose.

Statistic	Dose (Gy)
Mode	16.143
Expected	16.814
SD	1.778
95% CI LB	14.148
95% CI UB	19.953



# Partial body exposure

Higuera M, Puig P, Ainsbury EA, Vinnikov VA, Rothkamm, K (2015b).  
A New Bayesian Model Applied to Cytogenetic Partial Body  
Irradiation Estimation. *Radiat. Prot. Dosim.*, DOI: 10.1098/rspa.  
2014.0588.

# Bayes factor

The Bayesian alternative for PBI:

- Decision: does the sample come from a partial body exposure? In contrast to the frequentist  $u$ -test, the Bayarri *et al.* 2008 Bayes factor is proposed (ZIP vs. Poisson):

$$BF = \frac{n_0!}{(n+1)!} \sum_{j=0}^{n_0} \frac{(n-j)!}{(n_0-j)!} (1 - j/n)^{-(s+1/2)},$$

where  $n$ ,  $n_0$  and  $s$  are respectively the sample size and frequency of zeros, and the sum of the total number of chromosomal aberrations.

- Once the ZIP assumption is supported, a new Bayesian model is proposed for the dose and fraction of the body irradiated estimation.

# The model

The frequency of aberrations per cell is

$$Z \sim \text{ZIP} \left( \mu, \frac{1 - F}{F e^{-D/d_0} - F + 1} \right),$$

where analogously to Higuera *et al.* (2015a)

$$\mu|D \sim \text{Gamma} \left( \frac{f(D, \hat{\beta})^2}{v(D, \hat{\beta})}, \frac{f(D, \hat{\beta})}{v(D, \hat{\beta})} \right)$$

an application of Bayes' theorem shows the expression of the likelihood of  $\mathbf{D}$ ,  $\mathbf{F}$  and  $\mathbf{d}_0$  for the given test data

$$L(y|D, F, d_0) \propto (F e^{-D/d_0} - F + 1)^{-n} \sum_{j=1}^{n_0} \binom{n_0}{j} \frac{F^{n-j} (1 - F)^j}{(n - j)^s} P(X_j = s|D),$$

where  $\mathbf{X}_j$  is a random variable negative binomial distribution with mean and variance depending on  $\mathbf{j}$  and  $\mathbf{D}$ .

# The model

Considering  $\mathbf{D}$ ,  $\mathbf{F}$  and  $\mathbf{d}_0$  as independent random variables, and

$$D \sim \text{Gamma} \left( \frac{\hat{D}^2}{\hat{\sigma}_{\hat{D}}^2}, \frac{\hat{D}}{\hat{\sigma}_{\hat{D}}^2} \right); F \sim \mathcal{U}(0, 1); d_0 \sim \mathcal{U}(2.7, 3.5);$$

the joint posterior density,

$$P(D, F, d_0|y) = \frac{L(y|D, F, d_0)P(D, F, d_0)}{\int L(y|D, F, d_0)P(D, F, d_0)dDdFdd_0},$$

has a non-tractable form. The acceptance-rejection sampling is used to simulate the posterior distribution.

# Example

In a recent experiment to simulate PBI, unirradiated and irradiated blood at each dose was mixed. For instance, the exposed blood fraction comprised 10% for 2 Gy, whose distribution of dicentrics plus centric rings is

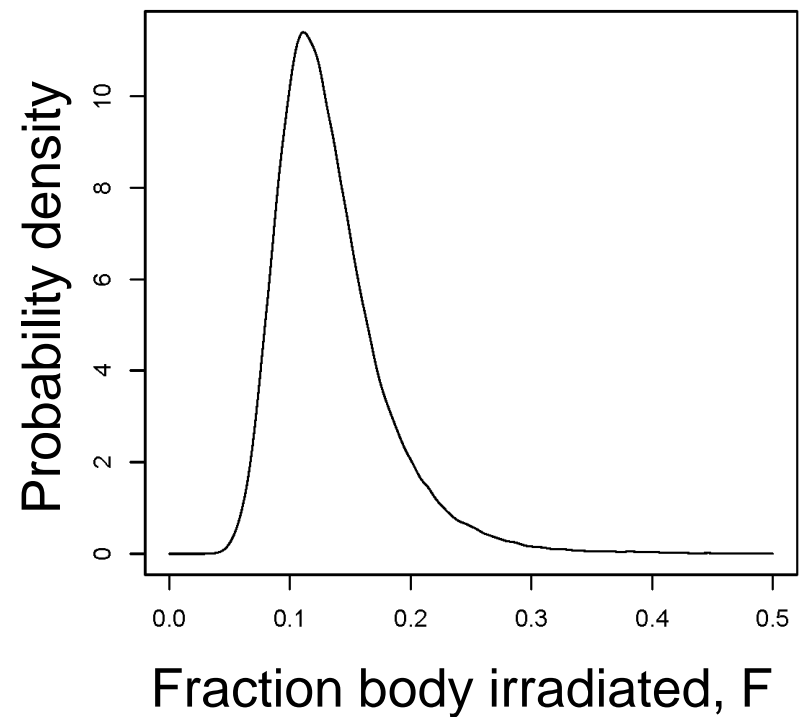
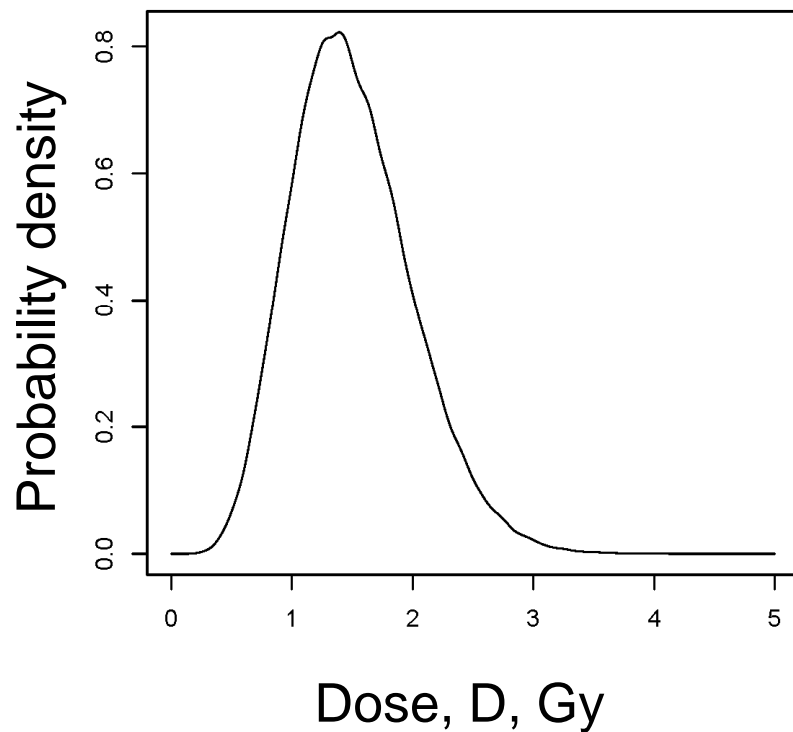
#Dic+CR	0	1	2
#cells	1043	16	3

The Bayes factor value for this sample gives ‘strong’ evidence in support to the ZIP assumption, because  $6 < 2 \log BF = 9.41 < 10$ .

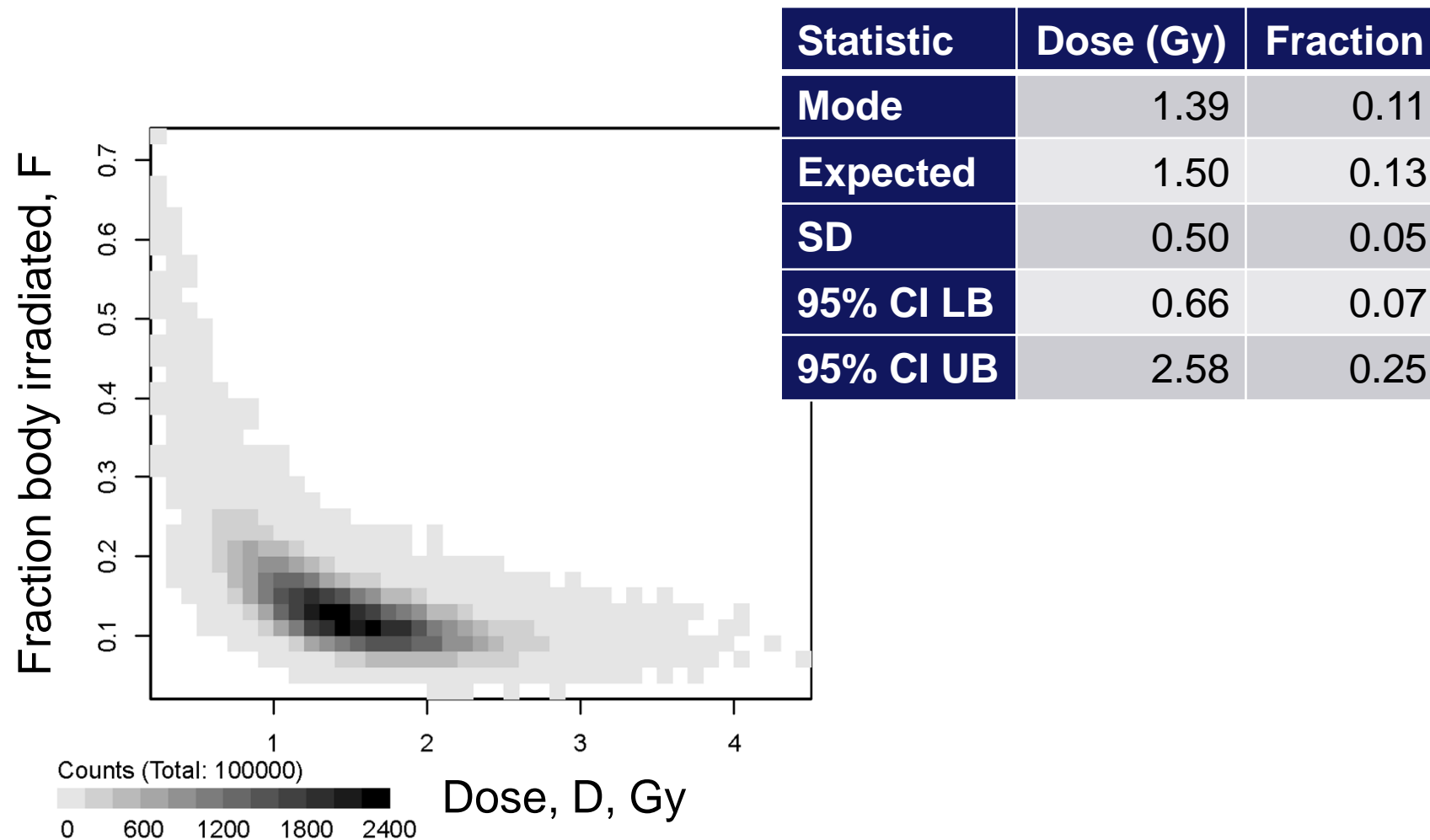
The calibration dose response data is taken from Vinnikov *et al.* 2013.



# Marginal posteriors



# Joint posterior and statistic summary



# Conclusions

- Novel solutions for statistical analysis of cytogenetic biological dosimetry data have been developed.
- New Bayesian models have been created and applied in practical cytogenetic dose estimation.
- Some of these models have been implemented in the R statistical software for biodosimetry laboratory researchers.
- These new solutions lead to more accurate quantification of statistical uncertainty associated with cytogenetic dose estimates.

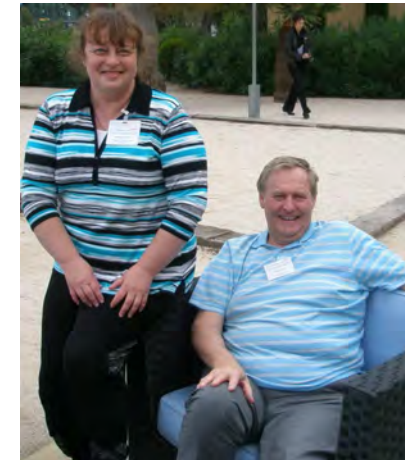
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# Acknowledgements



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**...and all the pieces matter.**

**Thanks for your attention!**